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# A Family of Estimators for Population Variance Using Two Auxiliary Variables

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#### **ABSTRACT**

This paper deals with the problem of estimating the population variance when some information on two auxiliary variables is available. It is shown that the proposed estimator is more efficient than the usual mean estimator and other existing estimators. The study is also extended to two-phase sampling. Theoretical results are supported by an empirical study.

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#### Introduction

In sample surveys, it is well known that the use of the auxiliary information increases the efficiency of the estimators. We can use the knowledge of the auxiliary variable x for estimating the study variable y. When the correlation between the study variate and the auxiliary variate is positive ratio method of estimation is most effective and when this correlation is negative product method of estimation can be employed effectively. Let x and z denote the study variates taking the values  $(x_i, z_i)$ .

In order to have an estimate of the study variable y. Isaki (1983) and Singh et al. (2007) proposed following estimators

$$t_1 = s_y^2 \left( \frac{S_x^2}{s_x^2} \right) \tag{1.1}$$

$$t_2 = s_y^2 \left( \frac{S_z^2}{s_z^2} \right) \tag{1.2}$$

$$t_{3} = s_{y}^{2} \exp\left(\frac{S_{x}^{2} - s_{x}^{2}}{S_{x}^{2} - s_{x}^{2}}\right)$$
(1.3)

$$t_4 = s_y^2 \exp\left(\frac{s_z^2 - S_z^2}{s_z^2 - S_z^2}\right)$$
 (1.4)

The Bias and MSE expression's of the estimator's  $t_i$  (i=1, 2, 3, 4) up to the first order of approximation are, respectively, given by

$$B(t_1) = \frac{S_y^2}{n} \left[ \delta_2 - \delta_3 \right] \tag{1.5}$$

$$\mathbf{B}(\mathbf{t}_2) = \frac{\mathbf{S}_y^2}{\mathbf{p}} \left[ \delta_4 - \delta_5 \right] \tag{1.6}$$

$$B(t_3) = \frac{S_y^2}{n} \left[ \frac{\delta_2}{8} - \frac{\delta_3}{2} + \frac{3}{8} \right]$$
 (1.7)

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$$B(t_4) = \frac{S_y^2}{n} \left[ \frac{\delta_4}{8} - \frac{\delta_5}{2} - \frac{5}{8} \right]$$
 (1.8)

$$MSE(t_1) = \frac{S_y^4}{n} [\delta_1 + \delta_2 - 2\delta_3]$$
 (1.9)

$$MSE(t_{2}) = \frac{S_{y}^{4}}{n} \left[ \delta_{1} + \delta_{4} - 2\delta_{5} \right]$$
 (1.10)

$$MSE(t_3) = \frac{S_y^4}{n} \left[ \delta_1 + \frac{\delta_2}{4} - \delta_3 - \frac{1}{4} \right]$$
 (1.11)

$$MSE(t_4) = \frac{S_y^4}{n} \left[ \delta_1 + \frac{\delta_4}{4} + \delta_5 - \frac{9}{4} \right]$$
 (1.12)

Following Naik and Gupta (1996) and Singh et al. (2007), we propose the estimators t<sub>5</sub> and t<sub>6</sub> as

$$t_{5} = s_{y}^{2} \left(\frac{S_{x}^{2}}{s_{x}^{2}}\right)^{\alpha_{1}} \left(\frac{S_{z}^{2}}{s_{z}^{2}}\right)^{\alpha_{2}}$$
(1.13)

$$t_{6} = s_{y}^{2} \exp\left(\frac{S_{x}^{2} - s_{x}^{2}}{S_{y}^{2} - s_{y}^{2}}\right)^{\beta_{1}} \exp\left(\frac{s_{z}^{2} - S_{z}^{2}}{s_{z}^{2} + S_{z}^{2}}\right)^{\beta_{2}}$$
(1.14)

where  $\alpha_1, \alpha_2, \beta_1$  and  $\beta_2$  are real constants.

## BIAS AND MSE of $t_5$ and $t_6$

To obtain the bias and MSE of  $t_5$  and  $t_6$  to the first degree of approximation, we define

$$s_y^2 = S_y^2 (1 + e_0), s_x^2 = S_x^2 (1 + e_1), s_z^2 = S_z^2 (1 + e_2)$$

Such that,  $E(e_i) = 0$ ; i = 0, 1, 2.

Also

$$E(e_0^2) = \frac{\left(\delta_1 - 1\right)}{n}, \ E(e_1^2) = \frac{\left(\delta_2 - 1\right)}{n}, \ E(e_2^2) = \frac{\left(\delta_4 - 1\right)}{n}$$

$$E(e_0e_1) = \frac{\left(\delta_3 - 1\right)}{n}, \ E(e_0e_2) = \frac{\left(\delta_5 - 1\right)}{n}, \ E(e_1e_2) = \frac{\left(\delta_6 - 1\right)}{n},$$

Expressing equation (1.13) in terms of e's, we have

$$\mathbf{t}_{5} = \mathbf{S}_{y}^{2} \left[ (1 + \mathbf{e}_{0}) (1 + \mathbf{e}_{1})^{-\alpha_{1}} (1 + \mathbf{e}_{2})^{-\alpha_{2}} \right] = \mathbf{S}_{y}^{2} \left[ 1 + \mathbf{e}_{0} - \alpha_{1} \mathbf{e}_{1} - \alpha_{2} \mathbf{e}_{2} + \frac{\alpha_{1} (\alpha_{1} + 1)}{2} \mathbf{e}_{1}^{2} + \frac{\alpha_{2} (\alpha_{2} + 1)}{2} \mathbf{e}_{2}^{2} + \alpha_{1} \alpha_{2} \mathbf{e}_{1} \mathbf{e}_{2} - \alpha_{1} \mathbf{e}_{0} \mathbf{e}_{1} - \alpha_{2} \mathbf{e}_{0} \mathbf{e}_{2} \right]$$

$$(2.1)$$

Subtracting  $\overline{Y}$  from both the sides of equation (2.1) and then taking expectation of both sides, we get the bias of the estimator  $t_5$ , up to the first order of approximation, as

$$B(t_{5}) = \frac{S_{y}^{2}}{n} \left[ \frac{\alpha_{1}(\alpha_{1}+1)}{2} (\delta_{2}-1) + \frac{\alpha_{2}(\alpha_{2}+1)}{2} (\delta_{4}-1) + \alpha_{1}\alpha_{2}(\delta_{6}-1) - \alpha_{1}(\delta_{3}-1) - \alpha_{2}(\delta_{5}-1) \right]$$
(2.2)

From (2.1), we have

$$(t_5 - \overline{Y}) = \overline{Y}[e_0 - \alpha_1 e_1 - \alpha_2 e_2]$$
(2.3)

Squaring both sides of (2.3) and then taking expectation, we get, MSE of the estimator  $t_5$ , up to the first order of approximation, as

$$MSE(t_5) = \frac{S_y^4}{n} \left[ (\delta_1 - 1) + \alpha_1^2 (\delta_2 - 1) + \alpha_2^2 (\delta_4 - 1) + 2\alpha_1 \alpha_2 (\delta_6 - 1) - 2\alpha_1 (\delta_3 - 1) - 2\alpha_2 (\delta_5 - 1) \right]$$
(2.4)

To obtain the bias and MSE of  $t_6$ , to the first order of approximation, we express equation (1.14) in term of e's, as

$$\begin{split} t_6 &= \overline{Y} \left( 1 + e_0 + \exp \left( \frac{-e_1}{2 + e_1} \right)^{\beta_1} \exp \left( \frac{e_2}{2 + e_2} \right)^{\beta_2} \right) \\ &= \overline{Y} \left[ 1 + e_0 - \frac{\beta_1 e_1}{2} + \frac{\beta_1^2 e_1^2}{4} + \frac{\beta_2 e_2}{2} - \frac{\beta_1 \beta_2 e_1 e_2}{4} + \frac{\beta_2^2 e_2^2}{4} + \frac{\beta_2 e_0 e_2}{2} - \frac{\beta_1 e_0 e_1}{2} \right] \end{split} \tag{2.5}$$

Subtracting  $\overline{Y}$  from both sides of equation (2.5) and then taking expectation of both sides, we get the bias of the estimator  $t_6$  up to the first order of approximation, as

$$B(t_6) = \frac{S_y^2}{n} \left\lceil \frac{\beta_1^2}{4} (\delta_2 - 1) + \frac{\beta_2^2}{4} (\delta_4 - 1) + \frac{\beta_1}{2} (\delta_3 - 1) - \frac{\beta_1 \beta_2}{2} (\delta_6 - 1) + \frac{\beta_2}{2} (\delta_5 - 1) \right\rceil$$
(2.6)

From (2.5), we have

$$(t_6 - \overline{Y}) = \overline{Y} \left[ e_0 - \frac{\beta_1 e_1}{2} + \frac{\beta_2 e_2}{2} \right]$$
 (2.7)

Squaring both sides of (2.7) and then taking expectation, we get the MSE of the estimator  $t_6$  up to the first order of approximation, as

$$MSE(t_6) = \frac{S_y^4}{n} \left[ (\delta_1 - 1) + \frac{\beta_1^2}{4} (\delta_2 - 1) + \frac{\beta_2^2}{4} (\delta_4 - 1) + \frac{\beta_1}{2} (\delta_3 - 1) - \frac{\beta_1 \beta_2}{2} (\delta_6 - 1) + \frac{\beta_2}{2} (\delta_5 - 1) \right]$$
(2.8)

#### 3. ANOTHER ESTIMATOR

Following Naik and Gupta (1996) and Singh et al. (2007), we propose another improved estimator t<sub>p</sub> as

$$t_{p} = w_{0}s_{y}^{2} + w_{1}s_{y}^{2} \left(\frac{S_{x}^{2}}{s_{x}^{2}}\right)^{\alpha_{1}} \left(\frac{S_{z}^{2}}{s_{z}^{2}}\right)^{\alpha_{2}} + w_{2}s_{y}^{2} exp \left(\frac{S_{x}^{2} - s_{x}^{2}}{S_{x}^{2} - s_{x}^{2}}\right)^{\beta_{1}} exp \left(\frac{s_{z}^{2} - S_{z}^{2}}{s_{z}^{2} + S_{z}^{2}}\right)^{\beta_{2}}$$
(3.1)

where  $\alpha_1, \alpha_2, \beta_1$  and  $\beta_2$  are real constants and  $w_i (i = 0,1,2)$  are suitably chosen constants whose values are to be determined later.

Expressing (3.1) in terms of e's, we have

$$t_{p} = \overline{Y}(1 + e_{0}) \left[ w_{0} + w_{1}(1 + e_{1})^{-\alpha_{1}} (1 + e_{2})^{-\alpha_{2}} + w_{2} \exp\left(\frac{-\beta_{1}e_{1}}{2}\right) \exp\left(\frac{\beta_{2}e_{2}}{2}\right) \right]$$
(3.2)

Expanding the right hand side of equation (3.2) and retaining terms, up to second power of e's, we have

$$\begin{split} t_{p} &= \overline{Y} \Bigg[ 1 + e_{0} + w_{1} \Bigg( \frac{\alpha_{1} (1 + \alpha_{1})}{2} e_{1}^{2} + \frac{\alpha_{2} (1 + \alpha_{2})}{2} e_{2}^{2} - \alpha_{1} e_{1} - \alpha_{2} e_{2} + \alpha_{1} \alpha_{2} e_{1} e_{2} - \alpha_{1} e_{0} e_{1} - \alpha_{2} e_{0} e_{2} \Bigg) \\ &+ w_{2} \Bigg( \frac{\beta_{1}^{2} e_{1}^{2}}{4} - \frac{\beta_{1} e_{1}}{2} - \frac{\beta_{1} e_{0} e_{1}}{2} - \frac{\beta_{1} \beta_{2} e_{1} e_{2}}{4} + \frac{\beta_{2}^{2} e_{2}^{2}}{4} + \frac{\beta_{2} e_{2}}{2} + \frac{\beta_{2} e_{0} e_{2}}{2} \Bigg) \Bigg]. \end{split} \tag{3.3}$$

subtracting  $\overline{Y}$  from both sides of (3.3) and then taking expectation of both sides, we get the bias of the estimator  $t_p$ , up to the first order of approximation as

$$\begin{split} B \Big( t_p \Big) &= w_1 \frac{S_y^2}{n} \Bigg[ \frac{\alpha_1 (1 + \alpha_1)}{2} (\delta_2 - 1) + \frac{\alpha_2 (1 + \alpha_2)}{2} (\delta_4 - 1) + \frac{\alpha_1}{2} (\delta_6 - 1) - \frac{\alpha_1 \alpha_2}{2} (\delta_6 - 1) + \frac{\alpha_2}{2} (\delta_5 - 1) \Bigg] \\ &+ w_2 \frac{S_y^2}{n} \Bigg[ \frac{\beta_1^2}{4} (\delta_2 - 1) + \frac{\beta_2^2}{4} (\delta_4 - 1) + \frac{\beta_1}{2} (\delta_6 - 1) - \frac{\beta_1 \beta_2}{2} (\delta_6 - 1) + \frac{\beta_2}{2} (\delta_5 - 1) \Bigg]. \end{split}$$

from (3.3), we have

$$(t_{p} - \overline{Y}) = \overline{Y} \left[ e_{0} + w_{1} \left( -\alpha_{1} e_{1} - \alpha_{2} e_{2} \right) + w_{2} \left( -\frac{\beta_{1} e_{1}}{2} + \frac{\beta_{2} e_{2}}{2} \right) \right]$$
(3.5)

Squaring both sides of (3.5) and then taking expectation, we get MSE of the estimator  $t_p$  up to the first order of approximation, as

$$MSE(t_p) = \frac{S_y^4}{n} \left[ (\delta_1 - 1) + w_1^2 P_1 + w_2^2 P_2 - 2w_1 P_3 - w_2 P_4 + w_1 w_2 P_5 \right]$$
(3.6)

Where, 
$$w_1 = \frac{4P_2P_3 - P_4P_5}{4P_1P_2 - P_5^2}$$

$$w_2 = \frac{2P_1P_4 - 2P_3P_5}{4P_1P_2 - P_5^2}$$

$$(3.7)$$

and

$$\begin{split} P_{1} &= \alpha_{1}^{2} \left(\delta_{2} - 1\right) + \alpha_{2}^{2} \left(\delta_{4} - 1\right) + 2\alpha_{1}\alpha_{2} \left(\delta_{6} - 1\right) \\ P_{2} &= \frac{1}{4} \left[\beta_{1}^{2} \left(\delta_{2} - 1\right) + \beta_{2}^{2} \left(\delta_{4} - 1\right) - 2\beta_{1}\beta_{2} \left(\delta_{6} - 1\right)\right] \\ P_{3} &= \alpha_{1} \left(\delta_{3} - 1\right) + \alpha_{2} \left(\delta_{5} - 1\right) \\ P_{4} &= \beta_{1} \left(\delta_{3} - 1\right) - \beta_{2} \left(\delta_{5} - 1\right) \\ P_{5} &= \alpha_{1}\beta_{1} \left(\delta_{2} - 1\right) - \alpha_{2}\beta_{2} \left(\delta_{4} - 1\right) + \alpha_{2}\beta_{1} \left(\delta_{6} - 1\right) - \alpha_{1}\beta_{2} \left(\delta_{6} - 1\right) \end{split}$$

#### **Double Sampling**

In certain practical situations when  $S_x^2$  is not known a priori, the technique of two-phase sampling is used. This scheme requires collection of information on x and z the first phase sample of size n' (n' < N) and on y for the second phase sample of size n (n < n') from the first phase sample.

The estimator's  $t_1$ ,  $t_2$ ,  $t_3$  and  $t_4$  in two-phase sampling take the following form

$$t_{d1} = s_y^2 \left( \frac{s_x'^2}{s_x^2} \right) \tag{4.1}$$

$$t_{d2} = s_y^2 \left( \frac{s_z'^2}{s_z^2} \right) \tag{4.2}$$

$$t_{d3} = s_y^2 \exp\left(\frac{s_x'^2 - s_x^2}{s_x'^2 - s_x^2}\right)$$
(4.3)

$$t_{d4} = s_y^2 \exp\left(\frac{s_z'^2 - S_z^2}{s_z'^2 - S_z^2}\right)$$
(4.4)

The bias and MSE expressions of the estimators  $t_{d1}$ ,  $t_{d2}$ ,  $t_{d3}$  and  $t_{d4}$  up to first order of approximation, are respectively given as

$$B(t_{d1}) = S_{v}^{2} [f_{1} \delta_{2} + f_{2} \delta_{3}]$$
(4.5)

$$B(t_{d2}) = \frac{S_y^2}{n'} \left[ \delta_4 - \delta_5 \right] \tag{4.6}$$

$$B(t_{d3}) = S_y^2 \left[ f_1 \frac{\delta_2}{8} + f_2 \frac{\delta_3}{2} \right]$$
 (4.7)

$$B(t_{d4}) = S_y^2 \left[ f_1 \frac{\delta_4}{8} + f_2 \frac{\delta_5}{2} \right]$$
 (4.8)

$$MSE(t_{d1}) = S_y^4 \left\lceil \frac{\delta_1 - 1}{n} + f_1(\delta_2 - 1) + 2f_2(\delta_3 - 1) \right\rceil$$
(4.9)

$$MSE(t_{d2}) = S_y^4 \left[ \frac{\delta_1 - 1}{n} + \frac{\delta_4 - 2\delta_5 + 1}{n'} \right]$$
(4.10)

$$MSE(t_{d3}) = S_y^4 \left[ \frac{\delta_1 - 1}{n} + \frac{f_1}{4} (\delta_2 - 1) + f_2(\delta_3 - 1) \right]$$
(4.11)

$$MSE(t_{d4}) = S_y^4 \left[ \frac{\delta_1 - 1}{n} + \frac{f_1}{4} (\delta_4 - 1) - f_2(\delta_5 - 1) \right]$$
(4.12)

where.

$$S_z^2 = \frac{1}{n'-1} \sum_{i=1}^n (z_i - \overline{z'})^2, \quad S_x'^2 = \frac{1}{n'-1} \sum_{i=1}^{n'} (x_i - \overline{x'})^2,$$

$$f_1 = \frac{1}{n} - \frac{1}{n'}, \quad f_2 = \frac{1}{n'} - \frac{1}{n}.$$

The estimator's t<sub>5</sub> and t<sub>6</sub>, in two phase sampling, takes the following form

$$t_{d5} = s_y^2 \left(\frac{s_x'^2}{s_x^2}\right)^{m_1} \left(\frac{S_z^2}{s_z'^2}\right)^{m_2}$$
(4.13)

$$t_{d6} = s_y^2 \exp\left(\frac{s_x'^2 - s_x^2}{s_x'^2 - s_x^2}\right)^{n_1} \exp\left(\frac{s_z'^2 - S_z^2}{s_z'^2 - S_z^2}\right)^{n_2}$$
(4.14)

where  $m_1$ ,  $m_2$ ,  $n_1$  and  $n_2$  are real constants.

# BIAS AND MSE OF $t_{d5}$ and $t_{d6}$

To obtain the bias and MSE of  $t_5$  and  $t_6$  to the first degree of approximation, we define

$$s_{y}^{2} = S_{y}^{2}(1 + e_{0}), s_{x}^{2} = S_{x}^{2}(1 + e_{1}^{2}), s_{z}^{2} = S_{z}^{2}(1 + e_{2}^{2})$$

Such that,  $E(e'_i) = 0$ ; i = 0, 1, 2.

Also,

$$E(e_0^2) = \frac{\left(\delta_1 - 1\right)}{n}, \ E(e_1'^2) = \frac{\left(\delta_2 - 1\right)}{n'}, \ E(e_2'^2) = \frac{\left(\delta_4 - 1\right)}{n'}$$

$$E(e_0e'_1) = \frac{(\delta_3 - 1)}{n'}, \ E(e_0e'_2) = \frac{(\delta_5 - 1)}{n'}, \ E(e'_1e'_2) = \frac{(\delta_6 - 1)}{n'},$$

Expressing equation (4.13) in terms of e's, we have

$$t_{d5} = \overline{Y}(1 + e_o)(1 + e_1)^{m_1}(1 + e_1)^{-m_1}(1 + e_2)^{-m_2}$$

Expanding the right hand side of above equation and retaining terms up to second power of e's, we have

$$t_{d5} = S_{y}^{2} \left[ 1 + e_{o} - m_{1}e_{1} + \frac{m_{1}(m_{1} + 1)}{2} e_{1}^{2} - m_{1}^{2}e_{1}e_{2} + \frac{m_{1}(m_{1} - 1)}{2} e_{1}^{'2} + \frac{m_{1}(m_{1} - 1)}{2} e_{1}^{'2} + \frac{m_{1}(m_{1} - 1)}{2} e_{2}^{'2} - m_{1}e_{1}e_{0} + m_{1}e_{0}e_{1} - m_{2}e_{0}e_{2}^{'2} \right]$$

$$(5.1)$$

Subtracting  $\overline{Y}$  from both sides of (5.1) and then taking expectation, we get the bias of the estimator  $t_{d5}$ , up to the first order of approximation, as

$$B(t_{d5}) = S_{y}^{2} \left\{ \frac{m_{1}(m_{1}+1)}{2} \frac{(\delta_{2}-1)}{n} - m_{1}^{2} \frac{(\delta_{6}-1)}{n'} + \frac{m_{1}(m_{1}-1)}{2} \frac{(\delta_{2}-1)}{n'} + \frac{m_{2}(m_{2}+1)}{2} \frac{(\delta_{4}-1)}{n'} - m_{1} \frac{(\delta_{3}-1)}{n} + m_{1} \frac{(\delta_{3}-1)}{n'} - m_{2} \frac{(\delta_{5}-1)}{n'} \right\}$$

$$(5.2)$$

From (5.1), we have

as

$$(t_{d5} - S_y^2) = S_y^2 \left[ e_0 - m_1 e_1 + m_2 e_1' - m_2 e_2' \right]$$
(5.3)

Squaring both sides of (5.3) and then taking expectations, we get MSE of the estimator t<sub>d5</sub>, up to the first order of approximation,

$$MSE(t_{d5}) = S_y^4 \left[ \frac{(\delta_1 - 1)}{n} + m_1^2 f_1(\delta_2 - 1) + m_2^2 \frac{(\delta_2 - 1)}{n'} + 2m_1 f_2(\delta_3 - 1) - 2m_2(\delta_5 - 1) \right]$$
(5.4)

Now to obtain the bias and MSE of t<sub>d6</sub> to the first order of approximation, we express equation (4.14) in terms of e's

$$t_{d6} = \overline{Y}(1 + e_0) \exp\left(\frac{n_1 e_1}{2}\right) \exp\left(\frac{-n_1 e_1}{2}\right) \exp\left(\frac{n_2 e_2}{2}\right)$$
(5.5)

Expanding the right hand side of equation (5.5) and retaining terms up to second power of e's, we have

$$t_{d6} = \overline{Y} \left( 1 + e_0 + \frac{n_1 e_1^2}{2} - \frac{n_1 e_1}{2} + \frac{n_1^2 e_1^2}{4} + \frac{n_2 e_2^2}{2} + \frac{n_2 e_2'^2}{4} + \frac{n_1 e_0 e_1'}{2} - \frac{n_1 e_0 e_1}{2} + \frac{n_2 e_0 e_2'}{2} \right)$$
(5.6)

subtracting  $\overline{Y}$  from both sides of (5.6) and then taking expectations, we get the bias of the estimator  $t_{d6}$  up to the first order of approximation, as

$$B(t_{d6}) = S_y^2 \left\{ \frac{n_1^2}{4} \frac{(\delta_2 - 1)}{n} + \frac{n_2^2}{4} \frac{(\delta_4 - 1)}{n'} + \frac{n_1}{2} f_2(\delta_3 - 1) + \frac{n_2}{2n'} (\delta_5 - 1) \right\}$$
(5.7)

From (5.6), we have

$$(t_{d6} - S_y^2) = S_y^2 \left( e_0 + \frac{n_1 e_1'}{2} - \frac{n_1 e_1}{2} + \frac{n_2 e_2'}{2} \right)$$
(5.8)

squaring both sides of (5.8) and then taking expectations, we get the MSE of  $t_{d6}$  up to the first order of approximation, as

$$MSE(t_{d6}) = S_y^4 \left[ \frac{(\delta_1 - 1)}{n} + \frac{n_1^2 f_1(\delta_2 - 1)}{4} + n_2^2 \frac{(\delta_4 - 1)}{n'} + 2n_1 f_2(\delta_3 - 1) + \frac{2n_2(\delta_5 - 1)}{n'} \right]$$
(5.9)

### IMPROVED ESTIMATOR tp IN TWO-PHASE SAMPLING

The estimator t<sub>p</sub> in double sampling is written as

$$t_{pd} = h_0 s_y^2 + s_y^2 \left(\frac{s_x'^2}{s_x^2}\right)^{m_1} \left(\frac{S_z^2}{s_z'^2}\right)^{m_2} + s_y^2 exp \left(\frac{s_x'^2 - s_x^2}{s_x'^2 - s_x^2}\right)^{n_1} exp \left(\frac{s_z'^2 - S_z^2}{s_z'^2 - S_z^2}\right)^{n_2}$$
(6.1)

Where,  $m_1, m_2, n_1$ , and  $n_2$  are real constants and  $h_i$  (i = 0,1,2) are suitably chosen constants whose values are to be determined later.

Expressing (6.1) in terms of e's, we have

$$t_{pd} = h_0 \overline{y} + h_1 \overline{y} (1 + e_1')^{m_1} (1 + e_1)^{-m_2} (1 + e_2')^{-m_2}$$

$$+ h_2 \overline{y} \exp\left(\frac{n_1 e_1'}{2}\right) \exp\left(\frac{-n_1 e_1}{2}\right) \exp\left(\frac{n_2 e_2'}{2}\right)$$
(6.2)

Expanding the right hand side of (6.2) and retaining terms up to second power of e's as

$$\begin{split} t_{pd} &= \overline{Y} \Bigg[ 1 + e_0 + h_1 \Bigg( \frac{m_1 \Big( m_1 + 1 \Big)}{2} e_1^2 - m_1 e_1 - m_1 e_0 e_1 - m_2^2 e_1 e_1 + \frac{m_1 \Big( m_1 - 1 \Big) 2}{2} e_1^2 + m_1 e_1 \\ &\qquad \qquad + m_1 e_0 e_1 - m_2 e_2 - m_2 e_0 e_2 + \frac{m_2 \Big( m_2 + 1 \Big)}{2} e_2^2 \Bigg) \\ &\qquad \qquad + h_2 \Bigg( \frac{n_1 e_1}{2} - \frac{n_1 e_1}{2} + \frac{n_1^2 e_1^2}{4} + \frac{n_2 e_2}{2} + \frac{n_2 e_2'^2}{4} + \frac{n_1 e_0 e_1'}{2} - \frac{n_1 e_0 e_1}{2} + \frac{n_2 e_0 e_2'}{2} \Bigg) \Bigg] \end{split} \tag{6.3}$$

Subtracting  $\overline{Y}$  from both the sides of (6.3) then taking expectations on both the sides, we get the bias of the estimator  $t_d$  up to the first order of approximation as

$$B(t_{pd}) = \overline{Y} \left[ h_1 f_3 C_{p_1}^2 \left( \frac{m_1^2}{2} + \frac{m_1}{2} + -m_1 K_{pb_1} \right) + h_1 f_2 C_{p_2}^2 \left( \frac{m_2^2}{2} + \frac{m_2}{2} - m_2 k_{pb_2} \right) \right.$$

$$\left. + h_3 f_3 C_{p_1}^2 \left( \frac{n_1^2}{8} + \frac{n_1}{8} - \frac{n_1}{2} K_{pb_1} \right) + h_3 f_2 C_{p_2}^2 \left( \frac{n_2^2}{8} + \frac{n_2}{8} - \frac{n_2}{2} k_{pb_2} \right) \right]$$

$$\left. - \frac{n_2 k_{pb_2}}{8} \right]$$

$$\left. + \frac{n_3 f_3 C_{p_1}^2 \left( \frac{n_1^2}{8} + \frac{n_1}{8} - \frac{n_1}{2} K_{pb_1} \right) + h_3 f_2 C_{p_2}^2 \left( \frac{n_2^2}{8} + \frac{n_2}{8} - \frac{n_2}{2} k_{pb_2} \right) \right]$$

$$\left. - \frac{n_2 k_{pb_2}}{8} \right)$$

From (6.3), we have

$$t_{pd} - \overline{Y} = \overline{Y} \left[ e_0 + h_1 \left( -m_1 e_1 + m_1 e_1 - m_2 e_2 \right) + h_2 \left( \frac{n_1 e_1}{2} - \frac{n_1 e_1}{2} + \frac{n_2 e_2}{2} \right) \right]$$
(6.5)

Squaring both sides of (6.5) and then taking expectations, we get MSE of the estimator  $t_p$  up to first order of approximation as

$$MSE(t_{pd}) = \frac{S_y^4}{n} [(\delta_1 - 1) + h_1^2 B_1 + h_2^2 B_2 - 2h_1 B_3 - h_2 B_4 + h_1 h_2 B_5]$$
(6.6)

where,

$$h_{1} = \frac{4B_{2}B_{3} - B_{4}B_{5}}{4B_{1}B_{2} - B_{5}^{2}}$$

$$h_{2} = \frac{2B_{1}B_{4} - 2B_{3}B_{5}}{4B_{1}B_{2} - B_{5}^{2}}$$
(6.7)

and

$$B_{1} = m_{2}^{2} \frac{(\delta_{4} - 1)}{n'} + m_{1}^{2} (\delta_{2} - 1) \left(\frac{1}{n} - \frac{1}{n'}\right)$$

$$B_{2} = \frac{1}{4} \left[ n_{2}^{2} \frac{(\delta_{4} - 1)}{n'} + n_{1}^{2} (\delta_{2} - 1) \left(\frac{1}{n} - \frac{1}{n'}\right) \right]$$

$$B_{3} = m_{1} \left(\frac{1}{n} - \frac{1}{n'}\right) (\delta_{2} - 1) + m_{2} \frac{(\delta_{4} - 1)}{n'}$$

$$B_{4} = n_{1} \left(\frac{1}{n} - \frac{1}{n'}\right) (\delta_{2} - 1) - n_{2} \frac{(\delta_{4} - 1)}{n'}$$

$$B_{5} = n_{1} m_{1} \left(\frac{1}{n} - \frac{1}{n'}\right) (\delta_{2} - 1) - n_{2} m_{2} \frac{(\delta_{4} - 1)}{n'}$$

#### **Empirical study**

To illustrate the performance of various estimators of  $S_v^2$ , we consider the data given in Murthy (1967, p.-226). The variates are

Y: output, x: number of workers, z: fixed capital

N=80, n'=25, n=10

$$\delta_1 = 2.2667 \ , \delta_2 = 3.65 \ , \ \delta_3 = 2.3377 \ , \ \delta_4 = 2.8664 \ , \ \delta_5 = 2.2208 \ , \ \delta_6 = 3.14$$

The percent relative efficiency (PRE) of various estimators of  $S_y^2$  with respect to conventional estimator  $S_y^2$  has been computed and displayed in the given table.

Table 7.1 PRE of different estimators of  $S_y^2$  with respect to  $S_y^2$ 

Choice of scalars								Т
$\mathbf{w}_0$	$\mathbf{w}_1$	w <sub>2</sub>	$\alpha_1$	$\alpha_2$	$\beta_1$	$\beta_2$	Estimator	PRE'S
1	0	0					s <sub>y</sub> <sup>2</sup>	100
0	1	0	1	0			t <sub>1</sub>	102.04
			0	1			t <sub>2</sub>	183.18
			-1	1			t <sub>5</sub>	25.61
0	0	1			1	0	t <sub>3</sub>	214.15
					0	1	t <sub>4</sub>	42.87
					1	-1	t <sub>6</sub>	104.78
$\mathbf{w}_0$	w <sub>1</sub>	w <sub>2</sub>	1	1	1	1	t <sub>p</sub>	282.39

In table 7.2 PRE of various estimators of in two-phase sampling are displayed.

Choice of scalars PRE'S **Estimator**  $\mathbf{h}_0$ h,  $n_1$  $h_1$  $m_2$  $m_1$  $n_2$  $s_y^2$ 1 0 0 100 0 0 0 1 1 1260.71  $t_{d1}$ 0 1 1033.90  $t_{d2}$ 1 1 1287.04  $t_{d5}$ 0 0 1270.76 1 1 0  $t_{d3}$ 0 1 513.85  $t_{d4}$ 1 1 1380.62  $t_{d6}$ 1 1 1 1472.25 1  $h_0$  $h_1$  $h_2$ t<sub>pd</sub>

Table 7.2: PRE of different estimators of  $S_{\nu}^2$  with respect to  $S_{\nu}^2$ 

#### **Conclusion:**

In this paper, we have suggested a class of estimators in single and double sampling for estimating population variance using two auxiliary variables. From Table 7.1 and Table 7.2, we infer that the proposed estimator  $t_p$  and its double sampling version  $t_{pd}$ , performs better than other estimators considered in this paper.

#### References

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